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Does Income Inequality Affect Capital Flows?

Evidence from Emerging Markets and Developing Economies

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Abstract

We assess the effect of income inequality on capital flows. We differentiate between aggregate capital inflows (external liabilities accumulation) and outflows (external assets accumulation) and disaggregated public and private capital inflows and outflows. We estimate dynamic panel data models using annual observations for Emerging Markets and Developing Economies during the 1999-2019 period. We find that the Top 1 and the Top 10 inequality measures are positive and statistically significant for aggregate and private inflows, and the Gini disposable income is statistically significant only for one explored method. The evidence also shows that there is a weak effect on private outflows, robust across methods only at the aggregate specification. The results also suggest that financial openness is positively associated with a greater effect of inequality.

Keywords: Income Inequality, Capital Flows, Financial Openness, Panel Data Models.

JEL Codes: D31, F21, F32, F41, C23.

I. Introduction

Income inequality could have a heterogeneous influence over the differential holding of external assets and external liabilities by distinct institutional sectors. Thus, it is not innocuous that external assets are in the hands of the public or private sector, nor that external liabilities belong to one or another type of economic agent. Ignoring this aspect of the analysis could lead to important consequences: for example, the valuation effect resulting from external shocks (i.e. an exchange rate depreciation) that affect the external assets and liabilities of the balance sheet of different economic agents, but that have different economic implications, particularly when we focus on the link between income inequality and these capital flows.

Pronounced inequality may be related to greater capital inflows and outflows due to: i) the higher propensity to accumulate external wealth showed by the superior deciles of the income distribution; ii) the external indebtedness (private capital inflows) held by the top incomes private sector, given their lower transactions costs and major facilities to access external capital markets; iii) the greater incentive to accumulate external public assets (government capital outflows) and allocate them in the form of Sovereign Wealth Funds (SWF) that could be employed to finance public transfers and reduce inequality; iv) the greater central banks' capital outflows that contribute to the increase of FX reserves that mitigate external vulnerability to global shocks and their associated regressive exchange rate depreciations; and v) the sizeable government capital inflows (public external liability) that could attend some negative shocks that increase inequality.

As far as we know, this topic has not been empirically examined. To accomplish this goal, we analyze how inequality affects capital flows at an aggregate (inflows and outflows) and at a disaggregated level (public and private inflows and outflows, separately), in both cases as a percentage of the GDP of each Emerging Market and Developing Economy (EMDE). We also study how income inequality affects net flows (i.e. capital outflows less capital inflows) at aggregate and disaggregate levels. Hence, our contribution is to provide aggregated and disaggregated dynamic panel data estimations using annual observations for 51, 38, and 35 EMDEs according to data availability for the 1999-2019 period. We estimate models which exclude and include the interaction between inequality indicators and financial openness. Increases in inequality and capital inflows and outflows at an aggregated or disaggregated level will depend on this interaction variable. We control for the usual variables that affect

capital flows, such as pull and push factors (Hannan, 2018), financial openness (Chinn Ito normalized index), and exchange rate regime (Ilzetzki et al. exchange rate regime updated classification).

We find heterogeneous effects between increases in inequality and capital flows not only at an aggregate level of analysis. We also report some novel empirical results considering the disaggregate level. The interaction between high inequality and preeminent financial openness explains the behavior of the public sector (which increases their public external liability, in terms of larger public capital inflows) and the conduct of the private sector too (which expands their private external liability). Public capital outflows could take the form of FX reserve accumulation by EMDEs' central banks, or an increase in SWF in the case of EMDEs' governments. Private capital outflows (private external assets) relate to a greater diversification of households and firms' portfolio decisions. These results intensify when increases in inequality interact with a larger financial openness.

Inequality and capital flows have important implications for macroprudential policies. EMDEs are exposed to capital flows that may trigger a balance of payment crisis, such as for instance shown in the prolific sudden stop literature. While several variables have been explored regarding the appropriate way of avoiding this, income inequality has not been considered yet. The empirical evidence here points out that changes in inequality measures trigger large capital flows.

The paper structure proceeds as follows. Section II assesses the related literature. Section III describes the database, the variables, and the information sources. Section IV presents the econometric model. Section V displays the main empirical results. Section VI offers some conclusions and economic policy implications.

II. Related literature

Our research finds support in different theoretical and empirical contributions related to four main topics: i) the positive association between income inequality and capital flows, mostly through their effects on the net foreign assets to GDP ratio (NFA/GDP, see Kumhof et al., 2012), ii) the positive correlation between unexpected shocks that increase inequality and the use of public debt (Carrera et al., 2023), iii) the utilization of FX reserves (Ortiz et al., 2017) or SWF to alleviate rises in inequality (Kemme et al.,

2021; Corneo, 2016), and iv) the direct relationship between financial openness, income inequality, and public indebtedness (Azzimonti et al., 2014).

Kumhof et al. (2012) built up a DSGE model with heterogeneous agents in an open economy setting, where the richest individuals receive dividends from firms and the poorest ones only earn wages. The former group is more likely to save and accumulate net external assets, while the latter is mostly paid in local currency. Thus, inequality affects the aggregate net external wealth position of a country. When bottom earners' income share declines at the expense of top earners, who have a much higher marginal propensity to save, top earners respond by increasing not only their consumption but also their desired wealth holdings. When an income shock primarily increases incomes derived from tradable assets, such as dividend incomes, actual wealth holdings increase by far more than desired wealth holdings, so that top earners borrow domestically and abroad, and the country runs a current account deficit. When the shock primarily increases incomes not derived from tradable assets, such as labor incomes, actual wealth holdings increase by far less than desired wealth holdings, top earners lend domestically and abroad, and the country runs a current account surplus.

Carrera et al. (2023) help us to understand why unexpected negative shocks to income inequality are usually associated with active fiscal responses from authorities, who prefer to finance progressive transfers with public borrowing in local and foreign currency instead of increasing taxes. These authors find that for EMDEs the interaction between the political cycle – proxied by the remaining time to complete the mandate – and income inequality is significant and positively related only to public debt, including the public external one. The marginal effect of inequality on the public debt is increasing in the share of the executive term completed. The empirical approach taken by these scholars used some arguments from Political Economy contributions to prove that policymakers frequently opt for using public debt to face unexpected shocks that increase income inequality and maximize their chances of being reelected.

According to Ortiz et al. (2017), FX reserves could be used to lessen the negative social consequences of negative shocks on income inequality through progressive fiscal transfers. In addition, it is important to bear in mind that less external vulnerability to global shocks is key to avoiding exchange rate pressures that increase income inequality. In that sense, Aizenman et al. (2023) document that an additional 10 percentage points of FX reserves/GDP held ex-ante were associated with 1.5 to 2 percent less exchange rate depreciation.

The link between increases in income inequality and SWF could be understood in terms of using these external funds with the aim of attending to social emergencies, like negative shocks that increase income inequality and could jeopardize political stability (Corneo, 2016). In that regard, Kemme et al. (2021) point out that income inequality and SWFs may be linked and influence FPI¹ flows. Each is thought to have a positive effect on FPI flows. However, if policymakers employ current assets and SWF revenue for domestic objectives, the converse may be true for the subgroup of economies that currently have substantial income inequality.

Last, but not least, public indebtedness in international markets has been boosted by a combination of deregulation of financial markets, financial innovation and sophistication, and international financial integration (Caballero and Krishnamurthy, 2009; Gourinchas and Rey, 2007; Lane and Milesi-Ferretti, 2007). Based on a multi-country model with incomplete markets and endogenous government borrowing, Azzimonti et al. (2014) show that governments choose higher levels of public debt if financial markets become internationally integrated and income inequality rises. Income inequality is associated with greater uninsurable idiosyncratic risks that result in a higher demand for safe assets and a lower interest rate, and consequently, higher government borrowing. So, governments might choose to incur higher levels of public debt in international financial markets that are deeper than domestic ones, particularly when a country becomes internationally integrated, and inequality increases.

These concurrent explanations are the first tentative to rationalize the relationship between inequality and capital flows using recent literature but clearly do not preclude any other explanation that should be tested in future works.

III. Database, variables, and information sources

The information sources are the databases on capital inflows and outflows by institutional sectors recently published by Avdjiev et al. (2022), the FRED St. Louis (VIX), and the World Bank (GDP growth rate). The inequality indicators (Top 1 and Top 10) come from the World Inequality Database (WID). In the case of the Gini disposable income, we use the Standardized World Income Inequality Database (SWIID). The additional control variables proceed from the Chinn Ito Database (financial openness

¹ Foreign Portfolio Investment.

index, normalized), and the Ilzetki et al. (2019) Database (exchange rate regime classification).

We employ Avdjiev et al. (2022) capital flows classification, so capital inflows are defined as liability flows, while capital outflows are defined as asset flows. The distinction between asset and liability flows allows liability flows to be interpreted as inflows from foreign agents and asset flows as outflows from domestic agents. This is the primary working definition of capital flows in BOP statistics and elsewhere, which we use for consistency across all data sources. These scholars identify capital flows in the domestic economy by source and destination sectors. The domestic economy refers to entities that are residents in that economy, according to a rule known as the “Residence Principle”, regardless of the entity’s nationality. This is the foundation from which the BOP data is compiled, which we compare when we perform our filling exercise. The term “sector” refers to institutional sectors such as general government, central banks, depository corporations other than the central bank (“banks”), and other sectors (“corporates”).

Table 1. Variables, definitions, and information sources

| Variables | Definitions | Information Sources |
|--|--|--|
| Capital Inflows (% of GDP) | Public Capital Inflows (in % of GDP) + Private Capital Inflows (in % of GDP) | Own elaboration based on data from “Gross Capital Flows by Banks Corporates and Sovereigns” (Avdjiev et al., 2022) |
| Capital Outflows (% of GDP) | Public Capital Outflows (in % of GDP) + Private Capital Outflows (in % of GDP) | Own elaboration based on data from “Gross Capital Flows by Banks Corporates and Sovereigns” (Avdjiev et al., 2022) |
| Public Capital Inflows (% of GDP) | Inflows Portfolio Debt Public + Inflows Other Investment Debt Public (Public = General Government + Central Bank) | Own elaboration based on data from “Gross Capital Flows by Banks Corporates and Sovereigns” (Avdjiev et al., 2022) |
| Public Capital Outflows (% of GDP) | Outflows Portfolio Debt Public + Outflows Other Investment Debt Public (Public = General Government + Central Bank) | Own elaboration based on data from “Gross Capital Flows by Banks Corporates and Sovereigns” (Avdjiev et al., 2022) |
| Private Capital Inflows (% of GDP) | Inflows Portfolio Debt Private + Inflows Other Investment Debt Private (Private = Depository Corporations, except the Central Bank + Other Sectors) | Own elaboration based on data from “Gross Capital Flows by Banks Corporates and Sovereigns” (Avdjiev et al., 2022) |
| Private Capital Outflows (% of GDP) | Outflows Portfolio Debt Private + Outflows Other Investment Debt Private (Private = Depository Corporations, except the Central Bank + Other Sectors) | Own elaboration based on data from “Gross Capital Flows by Banks Corporates and Sovereigns” (Avdjiev et al., 2022) |
| Top 1 | Pre-tax national income Top 1 share | World Inequality Database (WID) |
| Top 10 | Pre-tax national income Top 10 share | World Inequality Database (WID) |
| Gini disposable income | Coefficient Gini in household disposable (post-tax, post-transfer) | Standardized World Income Inequality Database (SWIID) |
| GDP growth rate | GDP real growth rate (annual %) | World Bank Database |
| VIX | Chicago Board Options Exchange Volatility Index | FRED St. Louis |
| Financial openness | Chinn-Ito index normalized (an increase in the index means a higher financial openness) | Chinn-Ito Database |
| Exchange rate regime (ERR) | A dummy variable (1 if ERR is fixed) | Ilzetzki et al. (2019) Database |

Table 2 shows the correlations between capital inflows and outflows as a percentage of GDP and three income inequality measures in some of these countries for the period 1999-2019.

Table 2. Correlations between capital flows (% of GDP) and income inequality

| Capital inflows (% of GDP) and income inequality | | | |
|---|--------------|----------------------|-----------------------|
| | <i>Total</i> | <i>Public Sector</i> | <i>Private sector</i> |
| <i>Top 1</i> | 0.078** | -0.001 | 0.088*** |
| <i>Top 10</i> | 0.068** | 0.017 | 0.070** |
| <i>Gini disposable income</i> | 0.032 | 0.034 | 0.023 |

| Capital outflows (% of GDP) and income inequality | | | |
|--|--------------|----------------------|-----------------------|
| | <i>Total</i> | <i>Public Sector</i> | <i>Private sector</i> |
| <i>Top 1</i> | 0.206*** | -0.030 | 0.207*** |
| <i>Top 10</i> | 0.121*** | 0.005 | 0.117*** |
| <i>Gini disposable income</i> | 0.007 | -0.005 | 0.008 |

Source: Own elaboration based on data from “Gross Capital Flows by Banks Corporates and Sovereigns” (Avdjiev et al., 2022), World Inequality Database (WID), and Standardized World Income Inequality Database (SWIID). ** p<0.05, *** p<0.01.

These correlations are statistically significant and positive in the case of capital inflows and outflows (total) and the case of private capital inflows and outflows. Conversely, they are not statistically significant in the case of the public sector capital flows. So, the private sector shows an interesting empirical association between capital inflows (increases in external private liabilities) and outflows (external private assets accumulation) and more regressive income distribution patterns.

These heterogeneous statistical preliminary results motivate us to estimate the effect of different income inequality measures on capital inflows and outflows at an aggregate level and then to provide a more granular picture of this relationship by examining the different institutional sector’s behavior (capital inflows and outflows’ reaction to

inequality increases) between the public sector and the private one when income disparities worsen.

IV. Econometric model

In order to study the relationship between aggregate and disaggregate capital flows as a proportion of GDP and different inequality indicators, we estimate an unbalanced dynamic panel data model with annual data for 1999-2019. The number of EMDEs used in the estimations varies between 51, 38, and 35, depending on data availability.²

We estimate a dynamic panel data model specification incorporating the dependent variable lagged one period. The autoregressive coefficient is significant and reflects some persistence in capital inflows and outflows to and from the public and private sectors as a proportion of GDP. In addition, inequality indexes fluctuate gradually over time since income distribution depends on the entire economic structure and shows high persistence. As a result, the data exploration confirms that the appropriate model for implementation is a dynamic autoregressive panel data one.

The specification of the dynamic panel data model equation is the following,

$$y_{i,t} = \alpha + \beta y_{i,t-1} + \gamma I_{i,t-1} + \delta x_{i,t-1} + \sigma I_{i,t-1} * FO_{i,t-1} + \zeta_t + \mu_i + \varepsilon_{i,t},$$

where $y_{i,t}$ denotes aggregate capital inflows and outflows, disaggregate public and private capital inflows and outflows, and net capital flows at aggregate and disaggregated levels, all of them expressed as a proportion of GDP, $y_{i,t-1}$ is the first lag

² The databases have the following countries:

51 EMDEs: Albania, Argentina, Bangladesh, Belarus, Bolivia, Brazil, Bulgaria, Chile, China, Colombia, Costa Rica, Cote d'Ivoire, Croatia, Dominican Republic, Ecuador, Egypt, Gabon, Ghana, Guatemala, Hungary, India, Indonesia, Jamaica, Jordan, Kazakhstan, Kenya, Lebanon, Macedonia, Malaysia, Mexico, Morocco, Nigeria, Pakistan, Papua New Guinea, Paraguay, Peru, Philippines, Poland, Romania, Russia, South Africa, Sri Lanka, Sudan, Thailand, Trinidad and Tobago, Tunisia, Turkey, Ukraine, Uruguay, Venezuela, and Vietnam.

38 EMDEs: Albania, Argentina, Bangladesh, Belarus, Bolivia, Brazil, Bulgaria, Chile, Colombia, Costa Rica, Croatia, Dominican Republic, Ecuador, Guatemala, Hungary, Jamaica, Jordan, Kazakhstan, Lebanon, Macedonia, Malaysia, Mexico, Nigeria, Pakistan, Papua New Guinea, Peru, Philippines, Poland, Romania, Russia, South Africa, Sri Lanka, Thailand, Trinidad and Tobago, Turkey, Ukraine, Uruguay, and Venezuela.

35 EMDEs: Albania, Argentina, Bangladesh, Belarus, Bolivia, Brazil, Bulgaria, Chile, Colombia, Costa Rica, Croatia, Dominican Republic, Ecuador, Guatemala, Hungary, Jamaica, Kazakhstan, Lebanon, Macedonia, Malaysia, Mexico, Nigeria, Pakistan, Peru, Philippines, Poland, Romania, Russia, South Africa, Sri Lanka, Thailand, Turkey, Ukraine, Uruguay, and Venezuela.

of each of these dependent variables, $I_{i,t-1}$ are the one period lagged inequality indicators (the Top 1, the Top 10, and the Gini disposable income), $x_{i,t-1}$ is the lagged one period vector of control variables used in Avdjiev et al. (2022: the logarithm of the VIX and the GDP growth rate, proxies of push and pull factors, respectively) plus four additional lagged one period control variables (a financial openness index, $FO_{i,t-1}$, a dummy variable for the exchange rate regime classification, an interaction variable, $I_{i,t-1} * FO_{i,t-1}$, which express the relationship between the income inequality indicators and the degree of financial openness, and an interaction variable between the logarithm of the VIX and the financial openness index, which measures how the transmission of push factors to EMDEs depends on the degree of liberalization of the financial account), ζ_t is a time-fixed effect, μ_i is a country fixed-effect, and $\varepsilon_{i,t}$ is the error term.

We estimate this equation by dynamic Fixed Effects (FE) and Bun and Kiviet's (2006, BK) correction for Nickell bias using the implementation for unbalanced panels of Bruno (2005). For this case, we cannot apply the usual System GMM estimators (Arellano and Bond, 1991; Blundell and Bond, 1998) because the number of instruments exceeds by far the number of EMDEs for all considered alternatives (see Roodman, 2009a,b).

Our interest lies in evaluating the statistical significance and the sign of the coefficients of the variables $I_{i,t-1}$, as well as the total effect linked not only to the impact of each income inequality measure on capital flows but also the interaction variable ($I_{i,t-1} * FO_{i,t-1}$), which can moderate or amplify the influence of the variables $I_{i,t-1}$ on the capital flows both at aggregate and disaggregate levels as a proportion of GDP.

One issue to be considered is that of endogeneity and potential biases when interpreting the coefficients in a causal fashion. Having access to international channels for capital flow transmissions may influence inequality, thus determining that both inequality and capital flows are simultaneously determined. While this is certainly a possibility we follow a causal interpretation whenever possible. First, we acknowledge our impossibility of using an identification mechanism, such that an instrumental variable, to isolate an exogenous effect on inequality. Second, we use the lagged values of inequality measures, such that current-year capital flows are not the determinant of that year's distribution of income. Third, we argue that inequality measures are

determined by income sources across heterogeneous individuals, and they do not necessarily relate to flows of already existing financial resources.

V. Results

Tables 3 and 5 (Tables 4 and 6) show the econometric estimations using the dynamic FE and BK methods, respectively. We consider inflows and outflows separately in Tables 3 and 4, and then net flows, calculated as the difference between outflows-inflows (see Tables 5 and 6).

The Top 1 and the Top 10 inequality measures are positive and statistically significant for aggregate and private inflows, and Gini disposable income only for the BK method. These results suggest that an increase in inequality attracts private capital inflows into EMDEs. We can also interpret this result as private agents taking external liabilities. To have an idea of its magnitude, suppose that the Top 1 of the richest population increases their participation in income by 1%, i.e. 0.01, then this would be associated with a long-term increment in capital inflows of $0.01 \times 36.27 / (1 - 0.33) = 0.54\%$ (from Table 3, column 1) or $0.01 \times 35.90 / (1 - 0.43) = 0.63\%$ (from Table 4, column 1) of the country GDP.

These results also suggest that there is a weak association with outflows, robust across methods only at the aggregate specification, but the evidence is not conclusive. In all cases, there is no effect on public inflows and outflows. Overall, this could be interpreted as the public sector being neutral to changes in inequality measures regarding its financing purposes.

Consider now the effects on the net flow variable. For this case, there appears a negative effect on total flows for the Top 1 and the Top 10, but a positive one for the Gini disposable income coefficient. This is not observed when disaggregating into private and public agents. The lack of a clear pattern also indicates that inflows and outflows correspond to different phenomena. That is, heterogeneous agents respond differently to the accumulation of external assets and liabilities.

When we add the interaction variable ($I_{i,t-1} * FO_{i,t-1}$) to the analysis (see Tables 7-10), in all cases the evidence suggests that financial openness is positively associated with a greater effect of inequality. The effects appear at the aggregate level, public inflows and outflows, and private outflows. Moreover, when considering the net flows, some

specifications indicate that inequality has a positive net effect. In contrast to the specifications without interactions, there are two main differences. First, capital outflows appear now as affected by inequality the greater the financial openness. This determines that restrictions force capital flows on capital mobility, and more unequal societies tend to increase capital outflows only when capital mobility is sufficiently high. Second, the results suggest that financial openness is necessary for the public channel relating inequality and capital flows (both in and out). This confirms Azzimonti et al. (2014) dynamics for public aggregates. That is, for either change in international reserves or SWE, there need to be specific financial conditions for inequality to play a role.

Table 3. Dynamic panel data model estimations through FE without interactions. EMDEs (1999-2019)

| | Capital Inflows (in % of GDP) | | | Capital Outflows (in % of GDP) | | | Public Capital Inflows (in % of GDP) | | | Public Capital Outflows (in % of GDP) | | | Private Capital Inflows (in % of GDP) | | | Private Capital Outflows (in % of GDP) | | |
|------------------------------|----------------------------------|---------|---------|-----------------------------------|---------|---------|---|---------|---------|--|--------|--------|--|---------|---------|---|---------|---------|
| | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) |
| Lag Top 1 | 36.27* | | | 18.13* | | | 5.14 | | | 0.08 | | | 32.79* | | | 19.20* | | |
| | (18.20) | | | (10.63) | | | (3.53) | | | (1.86) | | | (17.59) | | | (11.35) | | |
| Lag Top 10 | | 28.00* | | | -6.24 | | | 4.40 | | | -2.18 | | | 24.99* | | | -2.81 | |
| | | (14.01) | | | (13.69) | | | (2.73) | | | (2.03) | | | (14.26) | | | (11.17) | |
| Lag Gini Disposable | | | 14.64 | | -9.79 | | | -1.37 | | | 0.04 | | | 17.62 | | | -8.80 | |
| | | | (19.70) | | (25.62) | | | (3.20) | | | (3.01) | | | (20.56) | | | (25.86) | |
| Lag Capital Inflows | 0.33*** | 0.32*** | 0.34*** | | | | | | | | | | | | | | | |
| | (0.04) | (0.04) | (0.05) | | | | | | | | | | | | | | | |
| Lag Capital Outflows | | | | 0.22*** | 0.22*** | 0.22*** | | | | | | | | | | | | |
| | | | | (0.06) | (0.06) | (0.06) | | | | | | | | | | | | |
| Lag Public Capital Inflows | | | | | | | 0.51*** | 0.51*** | 0.51*** | | | | | | | | | |
| | | | | | | | (0.05) | (0.05) | (0.06) | | | | | | | | | |
| Lag Public Capital Outflows | | | | | | | | | | -0.14 | -0.14 | -0.14 | | | | | | |
| | | | | | | | | | | (0.12) | (0.12) | (0.12) | | | | | | |
| Lag Private Capital Inflows | | | | | | | | | | | | | 0.24*** | 0.23*** | 0.25*** | | | |
| | | | | | | | | | | | | | (0.07) | (0.07) | (0.07) | | | |
| Lag Private Capital Outflows | | | | | | | | | | | | | | | | 0.23*** | 0.24*** | 0.23*** |
| | | | | | | | | | | | | | | | | (0.05) | (0.05) | (0.05) |
| Observations | 859 | 859 | 831 | 362 | 362 | 355 | 859 | 859 | 831 | 362 | 362 | 355 | 863 | 863 | 834 | 364 | 364 | 357 |
| R-squared | 0.21 | 0.21 | 0.21 | 0.14 | 0.14 | 0.14 | 0.32 | 0.32 | 0.32 | 0.10 | 0.10 | 0.10 | 0.14 | 0.14 | 0.14 | 0.13 | 0.12 | 0.12 |
| Number of id | 51 | 51 | 51 | 38 | 38 | 35 | 51 | 51 | 51 | 38 | 38 | 35 | 51 | 51 | 51 | 38 | 38 | 35 |

Note: Robust standard errors in parentheses. *** p<0.01, ** p<0.05, * p<0.1. Additional controls not reported

Table 4. Dynamic panel data model estimations through BK without interactions. EMDEs (1999-2019)

| | Capital Inflows (in % of GDP) | | | Capital Outflows (in % of GDP) | | | Public Capital Inflows (in % of GDP) | | | Public Capital Outflows (in % of GDP) | | | Private Capital Inflows (in % of GDP) | | | Private Capital Outflows (in % of GDP) | | |
|------------------------------|----------------------------------|--------------------|--------------------|-----------------------------------|-------------------|-------------------|---|-------------------|--------------------|--|-----------------|-----------------|--|-------------------|-------------------|---|-------------------|-------------------|
| | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) |
| Lag Top 1 | 35.90*** (10.88) | | | 18.25*** (3.41) | | | 6.66 (5.25) | | | 0.73 (1.07) | | | 32.95*** (4.89) | | | 18.34 (15.97) | | |
| Lag Top 10 | | 29.64*** (2.68) | | | 3.09 (8.65) | | | 5.54*** (1.09) | | | -1.01 (1.67) | | 27.55*** (2.91) | | | 4.97 (19.64) | | |
| Lag Gini Disposable | | | 17.19*** (0.56) | | | 6.33 (44.79) | | | -4.32*** (0.38) | | | 1.57 (11.52) | | | 14.62 (10.11) | | | 5.58 (28.22) |
| Lag Capital Inflows | 0.43*** (0.03) | 0.42*** (0.03) | 0.44*** (0.05) | | | | | | | | | | | | | | | |
| Lag Capital Outflows | | | | 0.32*** (0.06) | 0.31*** (0.05) | 0.32*** (0.04) | | | | | | | | | | | | |
| Lag Public Capital Inflows | | | | | | | 0.51*** (0.03) | 0.51*** (0.03) | 0.51*** (0.04) | | | | | | | | | |
| Lag Public Capital Outflows | | | | | | | | | | -0.07 (0.06) | -0.07 (0.06) | -0.07 (0.07) | | | | | | |
| Lag Private Capital Inflows | | | | | | | | | | | | | 0.34*** (0.06) | 0.34*** (0.06) | 0.36*** (0.03) | | | |
| Lag Private Capital Outflows | | | | | | | | | | | | | | | | 0.33*** (0.01) | 0.33*** (0.01) | 0.33*** (0.07) |
| Observations | 859 | 859 | 831 | 362 | 362 | 355 | 859 | 859 | 831 | 362 | 362 | 355 | 863 | 863 | 834 | 364 | 364 | 357 |
| Number of id | 51 | 51 | 51 | 38 | 38 | 35 | 51 | 51 | 51 | 38 | 38 | 35 | 51 | 51 | 51 | 38 | 38 | 35 |

Note: Standard errors in parentheses. *** p<0.01, ** p<0.05, * p<0.1. Additional controls not reported

Table 5. Dynamic panel data model estimations through FE without interactions. EMDEs (1999-2019)

| | Net Capital Inflows (in % of GDP) | | | Net Public Capital Inflows (in % of GDP) | | | Net Private Capital Inflows (in % of GDP) | | |
|---------------------------------|--------------------------------------|-------------------|--------------------|---|-------------------|-------------------|--|------------------|------------------|
| | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) |
| Lag Top 1 | -23.54 (21.00) | | | -6.06 (6.75) | | | -14.38 (21.46) | | |
| Lag Top 10 | | -16.65 (17.89) | | | -6.14 (4.70) | | | -8.33 (16.28) | |
| Lag Gini Disposable | | | 30.84** (14.27) | | | 14.03* (7.43) | | | 12.74 (10.00) |
| Lag Net Capital Inflows | 0.11 (0.20) | 0.12 (0.19) | 0.12 (0.18) | | | | | | |
| Lag Net Public Capital Inflows | | | | 0.38*** (0.08) | 0.38*** (0.09) | 0.38*** (0.09) | | | |
| Lag Net Private Capital Inflows | | | | | | | 0.10 (0.11) | 0.11 (0.10) | 0.11 (0.10) |
| Observations | 362 | 362 | 355 | 362 | 362 | 355 | 364 | 364 | 357 |
| R-squared | 0.09 | 0.09 | 0.09 | 0.26 | 0.27 | 0.27 | 0.09 | 0.08 | 0.09 |
| Number of id | 38 | 38 | 35 | 38 | 38 | 35 | 38 | 38 | 35 |

Note: Robust standard errors in parentheses. *** p<0.01, ** p<0.05, * p<0.1. Additional controls not reported

Table 6. Dynamic panel data model estimations through BK without interactions. EMDEs (1999-2019)

| | Net Capital Inflows (in % of GDP) | | | Net Public Capital Inflows (in % of GDP) | | | Net Private Capital Inflows (in % of GDP) | | |
|---------------------------------|--------------------------------------|-------------------|-------------------|---|-------------------|-------------------|--|-------------------|------------------|
| | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) |
| Lag Top 1 | -24.83*** (2.09) | | | -5.17*** (0.42) | | | -16.39 (16.30) | | |
| Lag Top 10 | | -13.05 (8.34) | | | -1.84 (4.59) | | | -6.33 (17.75) | |
| Lag Gini Disposable | | | 34.74 (44.98) | | | 15.29 (20.58) | | | 17.30 (28.27) |
| Lag Net Capital Inflows | 0.19*** (0.05) | 0.20*** (0.05) | 0.20*** (0.06) | | | | | | |
| Lag Net Public Capital Inflows | | | | 0.50*** (0.03) | 0.49*** (0.03) | 0.50*** (0.04) | | | |
| Lag Net Private Capital Inflows | | | | | | | 0.19*** (0.00) | 0.19*** (0.01) | 0.19** (0.08) |
| Observations | 362 | 362 | 355 | 362 | 362 | 355 | 364 | 364 | 357 |
| Number of id | 38 | 38 | 35 | 38 | 38 | 35 | 38 | 38 | 35 |

Note: Standard errors in parentheses. *** p<0.01, ** p<0.05, * p<0.1. Additional controls not reported

Table 7. Dynamic panel data model estimations through FE with interactions. EMDEs (1999-2019)

| | Capital Inflows (in % of GDP) | | | Capital Outflows (in % of GDP) | | | Public Capital Inflows (in % of GDP) | | | Public Capital Outflows (in % of GDP) | | | Private Capital Inflows (in % of GDP) | | | Private Capital Outflows (in % of GDP) | | | | |
|---|----------------------------------|-------------------|-------------------|-----------------------------------|-------------------|--------------------|---|-------------------|-------------------|--|-----------------|-------------------|--|-------------------|-------------------|---|-----|-------------------|-------------------|-------------------|
| | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) | | |
| Lag Top 1 | 27.97 (20.38) | | | 0.37 (10.06) | | | 1.29 (3.80) | | | -3.83 (2.67) | | | 29.33 (20.64) | | | | | 4.91 (10.44) | | |
| Lag Top 10 | | 23.49 (15.07) | | | -14.93 (16.63) | | | 2.57 (2.77) | | | -4.00 (2.61) | | | 22.72 (15.84) | | | | -10.18 (13.81) | | |
| Lag Gini Disposable | | | 12.01 (21.86) | | | -20.59 (24.55) | | | -0.88 (3.38) | | | -2.77 (3.22) | | | 14.22 (22.60) | | | -17.64 (25.66) | | |
| Lag Top 1* Lag Financial Openness | 18.97 (16.56) | | | 30.74 (18.38) | | | 8.82* (5.14) | | | 6.76** (3.16) | | | 7.90 (13.89) | | | | | 24.71 (17.63) | | |
| Lag Top 10* Lag Financial Openness | | 12.33 (10.44) | | | 18.66 (11.96) | | | 5.02 (3.37) | | | 3.91 (2.35) | | | 6.16 (8.59) | | | | 15.70 (11.97) | | |
| Lag Gini Disposable* Lag Financial Openness | | | 7.30 (10.84) | | | 26.68** (13.05) | | | -1.38 (2.78) | | | 7.03*** (2.56) | | | 9.47 (10.44) | | | 22.00 (13.65) | | |
| Lag Capital Inflows | 0.33*** (0.04) | 0.32*** (0.04) | 0.34*** (0.05) | | | | | | | | | | | | | | | | | |
| Lag Capital Outflows | | | | 0.21*** (0.06) | 0.22*** (0.06) | 0.22*** (0.05) | | | | | | | | | | | | | | |
| Lag Public Capital Inflows | | | | | | | 0.51*** (0.06) | 0.51*** (0.05) | 0.51*** (0.06) | | | | | | | | | | | |
| Lag Public Capital Outflows | | | | | | | | | | -0.15 (0.12) | -0.15 (0.12) | -0.15 (0.12) | | | | | | | | |
| Lag Private Capital Inflows | | | | | | | | | | | | | 0.24*** (0.07) | 0.23*** (0.07) | 0.25*** (0.07) | | | | | |
| Lag Private Capital Outflows | | | | | | | | | | | | | | | | | | 0.23*** (0.05) | 0.23*** (0.05) | 0.23*** (0.05) |
| Observations | 859 | 859 | 831 | 362 | 362 | 355 | 859 | 859 | 831 | 362 | 362 | 355 | 863 | 863 | 834 | | | | | |
| R-squared | 0.21 | 0.21 | 0.21 | 0.15 | 0.14 | 0.14 | 0.32 | 0.32 | 0.32 | 0.10 | 0.10 | 0.11 | 0.14 | 0.14 | 0.14 | | | | | |
| Number of id | 51 | 51 | 51 | 38 | 38 | 35 | 51 | 51 | 51 | 38 | 38 | 35 | 51 | 51 | 51 | | | | | |

Note: Robust standard errors in parentheses. *** p<0.01, ** p<0.05, * p<0.1. Additional controls not reported

Table 8. Dynamic panel data model estimations through BK with interactions. EMDEs (1999-2019)

| | Capital Inflows (in % of GDP) | | | Capital Outflows (in % of GDP) | | | Public Capital Inflows (in % of GDP) | | | Public Capital Outflows (in % of GDP) | | | Private Capital Inflows (in % of GDP) | | | Private Capital Outflows (in % of GDP) | | |
|---|----------------------------------|--------------------|-------------------|-----------------------------------|--------------------|--------------------|---|-------------------|-------------------|--|-------------------|-------------------|--|--------------------|--------------------|---|-------------------|--------------------|
| | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) |
| Lag Top 1 | 29.29** (14.58) | | | 2.81 (8.07) | | | 2.70 (7.26) | | | -2.57 (1.73) | | | 31.44*** (1.06) | | | 5.57 (12.10) | | |
| Lag Top 10 | | 26.47*** (4.59) | | | -2.52 (5.85) | | | 3.71* (2.20) | | | -2.54** (1.24) | | | 26.84*** (9.69) | | 0.60 (16.61) | | |
| Lag Gini Disposable | | | 15.06 (12.81) | | | -3.40 (46.22) | | | -3.52 (6.08) | | | -1.01 (11.97) | | | 12.38*** (4.36) | | | -2.46 (28.79) |
| Lag Top 1* Lag Financial Openness | 15.91** (6.80) | | | 27.83*** (5.51) | | | 9.12** (3.63) | | | 5.94*** (0.82) | | | 4.31 (8.70) | | | 22.75*** (6.27) | | |
| Lag Top 10* Lag Financial Openness | | 9.67*** (2.99) | | | 12.96*** (4.92) | | | 5.21*** (1.81) | | | 3.39*** (0.75) | | | 3.01 (18.59) | | 10.28 (8.30) | | |
| Lag Gini Disposable* Lag Financial Openness | | | 6.66 (31.69) | | | 25.64*** (4.62) | | | -1.65 (16.39) | | | 6.72*** (1.13) | | | 7.28 (39.81) | | | 21.49*** (4.63) |
| Lag Capital Inflows | 0.43*** (0.03) | 0.42*** (0.03) | 0.44*** (0.05) | | | | | | | | | | | | | | | |
| Lag Capital Outflows | | | | 0.31*** (0.07) | 0.31*** (0.06) | 0.31*** (0.04) | | | | | | | | | | | | |
| Lag Public Capital Inflows | | | | | | | 0.51*** (0.03) | 0.51*** (0.04) | 0.51*** (0.04) | | | | | | | | | |
| Lag Public Capital Outflows | | | | | | | | | | -0.08 (0.07) | -0.08 (0.06) | -0.08 (0.06) | | | | | | |
| Lag Private Capital Inflows | | | | | | | | | | | | | 0.34*** (0.06) | 0.34*** (0.06) | 0.36*** (0.03) | | | |
| Lag Private Capital Outflows | | | | | | | | | | | | | | | | 0.33*** (0.01) | 0.32*** (0.01) | 0.33*** (0.07) |
| Observations | 859 | 859 | 831 | 362 | 362 | 355 | 859 | 859 | 831 | 362 | 362 | 355 | 863 | 863 | 834 | 364 | 364 | 357 |
| Number of id | 51 | 51 | 51 | 38 | 38 | 35 | 51 | 51 | 51 | 38 | 38 | 35 | 51 | 51 | 51 | 38 | 38 | 35 |

Note: Standard errors in parentheses. *** p<0.01, ** p<0.05, * p<0.1. Additional controls not reported

Table 9. Dynamic panel data model estimations through FE with interactions. EMDEs (1999-2019)

| | Net Capital Inflows (in % of GDP) | | | Net Public Capital Inflows (in % of GDP) | | | Net Private Capital Inflows (in % of GDP) | | |
|---|--------------------------------------|-------------------|-------------------|---|-------------------|-------------------|--|-------------------|-------------------|
| | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) |
| | Lag Top 1 | -22.11 (30.53) | | | -2.42 (8.55) | | | -18.40 (35.11) | |
| Lag Top 10 | | -12.00 (21.77) | | | -4.32 (4.98) | | | -5.95 (21.62) | |
| Lag Gini Disposable | | | 34.90* (18.12) | | | 12.07 (8.32) | | | 19.91* (11.58) |
| Lag Top 1* Lag Financial Openness | -2.48 (22.60) | | | -6.34 (8.17) | | | 6.91 (25.35) | | |
| Lag Top 10* Lag Financial Openness | | -9.99 (14.54) | | | -3.90 (5.58) | | | -5.06 (13.78) | |
| Lag Gini Disposable* Lag Financial Openness | | | -10.01 (15.48) | | | 4.94 (6.48) | | | -17.78 (12.01) |
| Lag Net Capital Inflows | 0.11 (0.20) | 0.12 (0.18) | 0.12 (0.18) | | | | | | |
| Lag Net Public Capital Inflows | | | | 0.38*** (0.08) | 0.38*** (0.09) | 0.38*** (0.09) | | | |
| Lag Net Private Capital Inflows | | | | | | | 0.10 (0.12) | 0.11 (0.10) | 0.10 (0.09) |
| Observations | 362 | 362 | 355 | 362 | 362 | 355 | 364 | 364 | 357 |
| R-squared | 0.09 | 0.09 | 0.09 | 0.27 | 0.27 | 0.27 | 0.09 | 0.08 | 0.09 |
| Number of id | 38 | 38 | 35 | 38 | 38 | 35 | 38 | 38 | 35 |

Note: Robust standard errors in parentheses. *** p<0.01, ** p<0.05, * p<0.1. Additional controls not reported

Table 10. Dynamic panel data model estimations through BK with interactions. EMDEs (1999-2019)

| | Net Capital Inflows (in % of GDP) | | | Net Public Capital Inflows (in % of GDP) | | | Net Private Capital Inflows (in % of GDP) | | |
|---|--------------------------------------|-------------------|--------------------|---|-------------------|-------------------|--|-------------------|---------------------|
| | (1) | (2) | (3) | (1) | (2) | (3) | (1) | (2) | (3) |
| Lag Top 1 | -25.11*** (7.68) | | | -2.12 (4.64) | | | -22.08* (12.18) | | |
| Lag Top 10 | | -7.68 (5.13) | | | 0.58 (2.62) | | | -3.28 (14.18) | |
| Lag Gini Disposable | | | 38.76 (46.03) | | | 14.18 (20.83) | | | 24.20 (29.52) |
| Lag Top 1* Lag Financial Openness | -0.08 (7.72) | | | -5.37 (5.16) | | | 9.23 (6.84) | | |
| Lag Top 10* Lag Financial Openness | | -11.08* (5.95) | | | -4.97 (3.38) | | | -6.07 (8.58) | |
| Lag Gini Disposable* Lag Financial Openness | | | -9.65*** (2.34) | | | 3.92*** (1.48) | | | -17.02*** (2.56) |
| Lag Net Capital Inflows | 0.19*** (0.06) | 0.20*** (0.05) | 0.20*** (0.06) | | | | | | |
| Lag Net Public Capital Inflows | | | | 0.50*** (0.03) | 0.49*** (0.02) | 0.50*** (0.04) | | | |
| Lag Net Private Capital Inflows | | | | | | | 0.18*** (0.00) | 0.19*** (0.00) | 0.19** (0.08) |
| Observations | 362 | 362 | 355 | 362 | 362 | 355 | 364 | 364 | 357 |
| Number of id | 38 | 38 | 35 | 38 | 38 | 35 | 38 | 38 | 35 |

Note: Standard errors in parentheses. *** p<0.01, ** p<0.05, * p<0.1. Additional controls not reported

VI. Conclusions

This paper provides empirical evidence of the relationship between increases in income inequality and its effects on capital flows. We examine not only the aggregate dimension of analysis (capital inflows and outflows) but also the disaggregated one (public and private sectors' capital inflows and outflows). The assessment of different institutional sectors' behavior is important to reveal a heterogeneous pattern that aggregate dimensions could hide.

A more unequal income distribution pattern in EMDEs is not innocuous regarding its effects on capital flows. EMDEs' institutional sectors react accordingly. The Top 1 and the Top 10 inequality measures are positive and statistically significant for aggregate and private inflows (i.e. increases in total and external private liabilities), and the Gini disposable income is statistically significant only for one explored method. These results suggest that an increase in inequality attracts private capital inflows into EMDEs. The evidence also shows that there is a weak effect on private outflows, robust across methods only at the aggregate specification.

When we add the interaction variable to the analysis, in all cases the results suggest that financial openness is positively associated with a greater effect of inequality. The effects appear at the aggregate level, public inflows and outflows, and private outflows. In contrast to the specifications without interactions, there are two main differences. First, capital outflows appear now as affected by inequality the greater the financial openness. This determines that restrictions force capital flows on capital mobility, and more unequal societies tend to increase capital outflows only when mobility is sufficiently high. Second, the evidence suggests that financial openness is necessary for the public channel connection between inequality and capital flows (both in and out).

These findings are important for policymakers to look out not only for political and social instability related to increases in income inequality but also for attending to possibly destabilizing macroeconomic effects linked to increments in both public and private external liabilities. The combination between a high stock of public external liability and a significant amount of external private assets could also have detrimental consequences for EMDEs' not only in terms of macroeconomic stability but also in further stock and flow inequality.

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